Errata for the ASM Study Manual for Exam P, Ninth Edition By Dr. Krzysztof M. Ostaszewski, FSA, CFA, MAAA Web site: http://www.krzysio.net E-mail: krzysio@krzysio.net

Posted September 21, 2011 The solution of Problem 29 in Practice Examination 19 should be: Solution.

The joint density, where positive, for 0 < x < 1 and 0 < y < x, is

$$f_{X,Y}(x,y) = f_Y(y|X=x) \cdot f_X(x) = \frac{1}{x} \cdot 2x = 2$$

The region where that joint density is positive is indicated with dotted lines in the graph below



Since the joint density is uniform, the conditional distribution of X given that Y = y is uniform on the range of values of x determined by the condition 0 < y < x < 1, i.e., the

interval (y,1), so that the variance is $\frac{(1-y)^2}{12}$. Answer E.

Posted August 31, 2011

The first sentence of the solution of Problem 14 in Practice Examination 18 should be:

Chi-square distribution is obtained as a sum of squares of independent identically distributed standard normal random variables, so we need to standardize these variables, add squares of those standardized variables, and hope we get one of the answers.

Posted August 6, 2011 In the third line of Problem 21 in Practice Examination 15, the words "and integer" should be "an integer".

Posted August 6, 2011

In the solution of Problem 6 in Practice Examination 9, in the first line, the words "for and" **should be** "for any".

Posted August 6, 2011

In the solution of Problem 1 in Practice Examination 9, the formula in the fifth line should be

$$\Pr(Y_{(1)} = 3) = \Pr(E - F) = \Pr(E) - \Pr(F).$$

instead of

$$\Pr(Y_{(1)}=3) = \Pr(F) - \Pr(F).$$

Posted August 4, 2011

The solution of Problem 7 in Practice Examination 13 should be rephrased as follows:

Solution.

The event of at least one color not being represented is the complement of the event of all three colors being represented, and all colors being represented simply means that we pick one red ball out of 3, one green ball out of 2, and one yellow ball out of 1. Thus

Pr(At least one color not drawn) =

$$= 1 - \Pr(\text{All colors drawn}) = 1 - \frac{\begin{pmatrix} 3 \\ 1 \end{pmatrix} \cdot \begin{pmatrix} 2 \\ 1 \end{pmatrix} \cdot \begin{pmatrix} 1 \\ 1 \end{pmatrix}}{\begin{pmatrix} 6 \\ 3 \end{pmatrix}} = 0.70.$$

You could also argue as follows. We are looking for the probability that all three balls are of the same color, or of two colors only. We have

$$\Pr(3 \text{ balls of one color}) = \Pr(3 \text{ red}) = \frac{1}{\binom{6}{3}} = \frac{3! \cdot 3!}{6!} = \frac{6}{4 \cdot 5 \cdot 6} = \frac{1}{20},$$

and

Pr(3 balls of two colors only) = Pr(2 red + 1 green) + Pr(2 red + 1 yellow) +

$$+ \Pr(2 \text{ green} + 1 \text{ red}) + \Pr(2 \text{ green} + 1 \text{ yellow}) = \\ = \frac{\binom{3}{2} \cdot \binom{2}{1}}{\binom{6}{3}} + \frac{\binom{3}{2} \cdot \binom{1}{1}}{\binom{6}{3}} + \frac{\binom{2}{2} \cdot \binom{3}{1}}{\binom{6}{3}} + \frac{\binom{2}{2} \cdot \binom{1}{1}}{\binom{6}{3}} = \frac{13}{20},$$

so that the total probability is $\frac{1}{20} + \frac{13}{20} = \frac{7}{10}$.

Answer B.

Posted August 3, 2011 The last sentence of Problem 11 in Practice Examination 16 should be: Calculate the variance of *Y* given that X > 3 and Y > 3. **instead of** Calculate the variance of *Y* given that and X > 3 and Y > 3.

Posted August 3, 2011 In Problem 7 in Practice Examination 14, the solution should start with the words Because you studied this manual instead of Because you studied his manual

Posted July 30, 2011 Problem 7 in Practice Examination 19 should start with

Let $X_1, X_2, ..., X_{36}$ and $Y_1, Y_2, ..., Y_{49}$ be independent random samples from distributions ... instead of

Let $x_1, x_2, ..., x_{36}$ and $y_1, y_2, ..., y_{49}$ be independent random samples from distributions ...

Posted July 30, 2011

The first line of the formula in the solution of Problem 6 in Practice Examination 19 should be

$$\Pr(2X - X^2 > 0) = \Pr(X(2 - X) > 0) = \Pr(X(X - 2) < 0) =$$

instead of

$$\Pr(2X - X^2 > 0) = \Pr(X(2 - X) > 0) = \Pr(X(X - 2) > 0) =$$

Posted July 28, 2011

In the formula in the solution of Problem 26 in Practice Examination 17, the formula should be

$$f_{X}(x) = F_{X}'(x) = -\frac{d}{dx} \sum_{k=0}^{3} \frac{x^{k} \cdot e^{-x}}{k!} = -\frac{d}{dx} \left(e^{-x} + xe^{-x} + \frac{1}{2}x^{2} \cdot e^{-x} + \frac{1}{6}x^{3} \cdot e^{-x} \right) = -\left(-e^{-x} + \left(e^{-x} - xe^{-x} \right) + \left(xe^{-x} - \frac{1}{2}x^{2} \cdot e^{-x} \right) + \left(\frac{1}{2}x^{2} \cdot e^{-x} - \frac{1}{6}x^{3} \cdot e^{-x} \right) \right) = \frac{1}{6}x^{3} \cdot e^{-x}.$$

The formula was missing a minus sign in the second line just after the first parenthesis. The rest of the solution is unaffected.

Posted July 28, 2011

In the solution of Problem 16 in Practice Examination 17, the word "bad" in the first sentence should be replaced by "bag".

Posted July 28, 2011

The second sentence of the solution of Problem 12 in Practice Examination 16 should be:

Box 1 contains 1 blue and 4 red marbles, box 2 contains 2 blue and 3 red marbles and box 3 contains 3 blue and 2 red marbles.

instead of

Box 1 contains 1 blue and 4 red marbles, box 2 contains 2 blue and 3 red marbles and box 3 contains 3 red and 2 blue marbles.

The rest of the solution is unaffected by this typo.

Posted July 26, 2011

In Problem 9 in Practice Examination 14, the third condition should be: (ii) The future lifetimes follow a Weibull distribution with $\alpha = 1.5$ and $\beta = 2.0$ for smokers, and $\alpha = 2.0$ and $\beta = 2.0$ for nonsmokers.

Also, the survival function of the Weibull distribution should be given as

 $s_T(t) = e^{-\left(\frac{t}{\alpha}\right)^{\beta}}.$

Posted June 26, 2011

In Problem 29 in Practice Examination 19, the last sentence should be: Find the variance of the conditional distribution of X, given Y = y.

Posted March 10, 2011

The second sentence of the solution of Problem 10 in Practice Examination 1 should be:

As the policy has a deductible of 1 (thousand), the claim payment is

$$Y = \begin{cases} 0, & \text{when there is no damage, with probability 0.94,} \\ \max(0, X - 1), & \text{when } 0 < X < 15, \text{ with probability 0.04,} \\ 14, & \text{in the case of total loss, with probability 0.02.} \end{cases}$$

Posted January 25, 2011

The last two sentences of the solution of Problem 11 in Practice Examination 16 should be replaced by

But the memoryless property of the exponential distribution tells us that *Y* and (Y-3|Y>3) have the same distribution. Note, however, that

$$(Y|Y>3) = 3 + (Y-3|Y>3),$$

so that

$$\operatorname{Var}(Y|Y>3) = \operatorname{Var}(Y-3|Y>3) = \operatorname{Var}(Y).$$

This implies that

$$\operatorname{Var}(Y|\{X>3\} \cap \{Y>3\}) = \operatorname{Var}(Y|Y>3) = \operatorname{Var}(Y) = \frac{1}{2^2} = 0.25.$$

Answer A.

Posted January 15, 2011

In Problem 16 in Practice Examination 6, the calculation of the second moment of *X* **should be:**

$$E(X^2) = \frac{1}{4} \cdot 0^2 + \frac{3}{4} \cdot \underbrace{(1+1)}_{\text{Second moment of }T} = \frac{3}{2}.$$

instead of

$$E(X^2) = E(X) = \frac{1}{4} \cdot 0^2 + \frac{3}{4} \cdot \underbrace{(1+1)}_{\text{Second moment of } T} = \frac{3}{2}$$

Posted July 24, 2010

In the solution of Problem 21 in Practice Examination 6, the statement under the first expression on the right-hand side of the third to last formula should be:

number of ways to		number of ways to
pick ordered samples	instead of	pick ordered samples
of size 2 from	mottau or	of size $n-2$ from
population of size n		population of size n

Posted July 17, 2010 Practice Examinations: An Introduction on page 109, the third sentence of the last section should be:

Practice examinations 6-18 are meant to be more challenging.



Some numbers in the diagram were mistyped.

Posted January 5, 2010 In the description of the gamma distribution in Section 2, the condition for the range of its MGF should be $t < \beta$, not $0 < t < \beta$.

Posted January 1, 2010

The Course P/1 syllabus updated for 2010 no longer contains direct references to chi-square, beta, Pareto, Weibull, and lognormal distributions. My interpretation of this change is that you do not need to memorize the details of chi-square, beta, Pareto and Weibull distributions, but you still should familiarize yourself with them. Since lognormal has a direct connection to normal, I think you should know that connection.

Posted November 20, 2009

Answers A and B in Problem 10, Practice Examination 9, have the symbol τ mistyped as *r* in the numerator, and they should be:

A.
$$f_Y(y) = \frac{\tau \theta y^{\tau-1}}{(y+\theta)^{\tau+1}}$$
 B. $f_Y(y) = \frac{\alpha \theta^{\alpha} \tau y^{\tau-1}}{(y^{\tau}+\theta)^{\alpha+1}}$

Posted November 17, 2009

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In Problem 17 in Practice Examination 14, answer choice A should be

$$f_{Y}(y) = \begin{cases} 0 & y < 0, \\ e^{1-e^{2}} \left(e^{ey} + e^{-ey} \right) & 0 < y < e, \\ e^{1-e^{2}} \cdot e^{-ey} & y \ge e. \end{cases}$$

and answer choice D should be:

$$f_{Y}(y) = \begin{cases} 0 & y < 0, \\ e^{e^{2}-1}(e^{ey} + e^{-ey}) & 0 < y < e, \\ e^{e^{2}-1} \cdot e^{-ey} & y \ge e. \end{cases}$$

The last sentence of the solution should be: Therefore, we can take

$$f_{Y}(y) = \begin{cases} 0 & y < 0, \\ e^{1-e^{2}} \left(e^{ey} + e^{-ey} \right) & 0 < y < e, \\ e^{1-e^{2}} \cdot e^{-ey} & y \ge e. \end{cases}$$

Posted September 2, 2009 In Problem 17 in Practice Examination 17, this statement and $f_N(n+1) > f_N(n)$ for n = 0, 1, 2, 3, 4. **should be removed.**

Posted July 1, 2009 In Section 2, the general definition of a percentile should be the 100-*p*-th *percentile* of the distribution of X is the number x_p which satisfies both of

the following inequalities: $\Pr(X \le x_p) \ge p$ and $\Pr(X \ge x_p) \ge 1 - p$.

Posted June 19, 2009

In the discussion of the mode of Poisson distribution on page 45, the expression and as soon $n \ge \lambda - 1$,

should be:

and as soon as $n \ge \lambda - 1$,

Posted June 18, 2009

The last formula in the solution of Problem 19 of Practice Examination 5 should be $\Pr(X \ge 10) \le \frac{1}{4^2} = \frac{1}{16}$ instead of $\Pr(X \ge 10) < \frac{1}{4^2} = \frac{1}{16}$.

Posted June 17, 2009

In the statement of Problem 9 in Practice Examination 7, Pr(X > 800) should be $Pr(X \ge 800)$, and in the solution, all inequalities should be changed accordingly.

Posted May 17, 2009

Problem 25 in Practice Examination 18 and its solution should be:

P Sample Exam Questions, Problem No. 128, also Dr. Ostaszewski's online exercise posted November 15, 2008

An insurance agent offers his clients auto insurance, homeowners insurance and renters insurance. The purchase of homeowners insurance and the purchase of renters insurance are mutually exclusive. The profile of the agent's clients is as follows:

i) 17% of the clients have none of these three products.

ii) 64% of the clients have auto insurance.

iii) Twice as many of the clients have homeowners insurance as have renters insurance.

iv) 35% of the clients have two of these three products.

v) 11% of the clients have homeowners insurance, but not auto insurance.

Calculate the percentage of the agent's clients that have both auto and renters insurance.

A. 7% B. 10% C. 16% D. 25% E. 28%

Solution.

Let *H* be the event of a client having a homeowners insurance, *R* be the event of a client having a renters insurance, and *A* be the event of a client having an auto insurance. We are given that $\Pr((H \cup R \cup A)^{c}) = 0.17$, so that $\Pr(H \cup R \cup A) = 0.83$. We are also given that $H \cap R = \emptyset$, so that $\Pr(H \cap R) = 0$. Note that $A \cap H \cap R \subset H \cap R$, so that $\Pr(A \cap H \cap R) = 0$, as well. Also, $\Pr(A) = 0.64$, $\Pr(H) = 2\Pr(R)$, and $\Pr(A \cap H) - \Pr(A \cap H \cap R) + \Pr(A \cap R) - \Pr(A \cap H \cap R) = \Pr(A \cap H) + \Pr(A \cap R) = 0.35$. Finally, we are also given that $\Pr(H - A) = \Pr(H - (A \cap H)) = \Pr(H) - \Pr(A \cap H) = 0.11$. The quantity we are looking for is $\Pr(A \cap R)$. First note that $0.83 = \Pr(H \cup R \cup A) = \Pr(H) + \Pr(R) + \Pr(R) + \Pr(A) - \Pr(H \cap R) - \Pr(H \cap A) = \Pr(H \cap A) - \Pr(R \cap A) + \Pr(R \cap A) = 3\Pr(R) + 0.64 - 0.35$

This means that Pr(R) = 0.18 and consequently Pr(H) = 0.36. But

 $0.36 = \Pr(H) = \Pr(H \cap A) + \Pr(H - A) = \Pr(H \cap A) + 0.11,$

so that $\Pr(H \cap A) = 0.25$ and

 $\Pr(A \cap R) = 0.35 - 0.25 = 0.10.$

Answer B.

Posted April 23, 2009

In Problem 20, Practice Examination 14, the first sentence of the solution should be: Let us define the following random variables:

X: time until death of Dwizeel by causes other than their private plane crash,

Y: time until death of Satellite Component by causes other than their private plane crash, Z: time until death of Dwizeel and Satellite Component as a result of their private plane crash.

The solution has the words "of" mistyped as "od."

Posted April 23, 2009 In Problem 9, Practice Examination 14, the Greek letter τ in the statement of the problem should be replaced by the Greek letter α .

Posted April 6, 2009 In the text of Problem 4 in Practice Examination 11, the sentence: We put that chip aside and pick a second chip from the same contained. **should be:** We put that chip aside and pick a second chip from the same container.

Posted April 6, 2009

In the text of Problem 7 in Practice Examination 3, the word "whwther" should be "whether".

Posted March 19, 2009

The solution of Problem 3 in Practice Examination 8 should be:

Let E be the event that a new insured is accident-free during the second policy year, and F be the event that a new insured is accident-free during the first policy year, and let G be the event that this new insured was accident-free the last year, before the policy was issued. Note that for any year only the previous year affects a given year, but not the year before that. Therefore

$$\Pr(E) = \Pr((E \cap F \cap G) \cup (E \cap F \cap G^{C}) \cup (E \cap F^{C} \cap G) \cup (E \cap F \cap G^{C})) =$$

$$= \Pr(E \cap F \cap G) + \Pr(E \cap F \cap G^{C}) + \Pr(E \cap F^{C} \cap G) + \Pr(E \cap F^{C} \cap G^{C}) =$$

$$= \Pr(G) \cdot \Pr(F|G) \cdot \Pr(E|F \cap G) + \Pr(G^{C}) \cdot \Pr(F|G^{C}) \cdot \Pr(E|F \cap G^{C}) +$$

$$+ \Pr(G) \cdot \Pr(F^{C}|G) \cdot \Pr(E|F^{C} \cap G) + \Pr(G^{C}) \cdot \Pr(F^{C}|G^{C}) \cdot \Pr(E|F^{C} \cap G^{C}) =$$

$$= 0.7 \cdot 0.8 \cdot 0.8 + 0.3 \cdot 0.6 \cdot 0.8 + 0.7 \cdot (1 - 0.8) \cdot 0.6 + 0.3 \cdot (1 - 0.6) \cdot 0.6 = 0.748.$$

Answer E.

Posted March 5, 2009

The first sentence of Problem 16 in Practice Examination 6 should end with t < 1, instead of t > 1.

Posted March 2, 2009

The properties of the cumulant moment-generating function should be: The cumulant generating function has the following properties:

$$\begin{split} \psi_{X}(0) &= 0, \qquad \qquad \frac{d}{dt} \ln E\left(e^{tX}\right)\Big|_{t=0} = \frac{E\left(Xe^{tX}\right)}{E\left(e^{tX}\right)}\Big|_{t=0} = E(X), \\ \frac{d^{2}}{dt^{2}}\psi_{X}(t)\Big|_{t=0} &= \frac{d}{dt} \frac{E\left(Xe^{tX}\right)}{E\left(e^{tX}\right)}\Big|_{t=0} = \frac{E\left(X^{2}e^{tX}\right)E\left(e^{tX}\right) - E\left(Xe^{tX}\right)E\left(Xe^{tX}\right)}{\left(E\left(e^{tX}\right)\right)^{2}}\Big|_{t=0} = \operatorname{Var}(X), \\ \frac{d^{3}}{dt^{3}}\psi_{X}(t)\Big|_{t=0} &= E\left(\left(X - E(X)\right)^{3}\right), \\ \text{for } k > 3, \end{split}$$

but

$$\left. \frac{d^k}{dt^k} \boldsymbol{\psi}_X(t) \right|_{t=0} = \boldsymbol{\psi}_X^{(k)}(0) < E\Big(\big(X - E(X)\big)^k\Big).$$

Also, if X and Y are independent (we will discuss this concept later), $\psi_{aX+b}(t) = \psi_X(at) + bt$, and $\psi_{X+Y}(t) = \psi_X(t) + \psi_Y(t)$.

Posted January 13, 2009

In Practice Examination 8, Problem 24, the calculation of the expected value had a typo, an extra, unnecessary p in the second line, and it instead should be:

$$E(X) = \sum_{k=1}^{+\infty} k \cdot \Pr(X = k) = 1 \cdot \Pr(X = 1) + \sum_{k=2}^{+\infty} k \cdot \Pr(X = k) =$$

$$= p + \sum_{k=1}^{+\infty} (k+1) \cdot \Pr(X = k+1) = p + \sum_{k=1}^{+\infty} k \cdot \Pr(X = k+1) + \sum_{k=1}^{+\infty} \Pr(X = k+1) =$$

$$= p + \sum_{k=1}^{+\infty} k \cdot \underbrace{(1-p) \cdot \Pr(X = k)}_{=\Pr(X=k+1)} + \underbrace{\left(\underbrace{\sum_{k=0}^{+\infty} \Pr(X = k+1)}_{\text{this sum is equal to 1}} \right) - \Pr(X = 0+1)}_{\text{this sum is equal to 1}} =$$

$$= p + (1-p) \cdot \sum_{k=1}^{+\infty} k \cdot \Pr(X = k) + (1 - \Pr(X = 1)) =$$

$$= p + (1-p) \cdot E(X) + (1-p) = 1 + (1-p) \cdot E(X).$$

Posted November 8, 2008

Problem 11 in Practice Examination 13 had several typos, and it should be:

Mr. Warrick Beige is gambling at the newly opened You Was Robbed casino in Nevada. In the game he is playing, first he has to choose one of two coins: coin A or coin B. Both coins are unfair. Coin A has the probability of heads of 0.60, and coin B has the probability of heads of 0.40. Mr. Beige pays \$20 to enter the game. He chooses a coin

randomly, but the chances of picking the coins are not equal. He has 40% chance of picking coin A and 60% chance of picking coin B. Then he tosses the coin chosen. If the result is heads, he is paid \$250. If the result is tails, he pays \$200. For an additional payment of x dollars, Mr. Beige can test a coin chosen: he can toss it once and based on the result, either walk away and get \$20 paid initially back (he will do this if the first toss results in tails), or toss the same coin again (he will do this if the first toss results in heads). Assuming that Mr. Beige values all gambles based on the expected value of the payoff (i.e., he is *risk-neutral*), calculate the value of x such that Mr. Beige is indifferent between testing a coin and not testing it.

A. \$6.40 B. \$4.00 C. \$2.50 D. \$0.00 E. -\$2.00

Solution.

We begin by labelling the events:

- A: Coin A is picked,
- B: Coin B is picked,
- H_1 : First toss results in heads,
- T_1 : First toss results in tails,
- H_2 : Second toss results in heads,
- T_2 : Second toss results in tails.

We know that Pr(A) = 0.40, Pr(B) = 0.60, $Pr(H_1|A) = 0.60$, and $Pr(H_1|B) = 0.40$. Therefore, the probability of getting heads in the first toss is

 $\Pr(H_1) = \Pr(H_1|A) \cdot \Pr(A) + \Pr(H_1|B) \cdot \Pr(B) =$

 $= 0.40 \cdot 0.60 + 0.60 \cdot 0.40 = 0.24 + 0.24 = 0.48.$

Therefore, Mr. Beige's expected gain on this game without testing first is

 $-\$20 + 0.48 \cdot \$250 + 0.52 \cdot (-\$200) = -\$4.00.$

Note that the two coin tosses are independent, and hence

$$\Pr(H_2|H_1) = \frac{\Pr(H_1 \cap H_2)}{\Pr(H_1)} = \frac{\Pr(H_1 \cap H_2|A) \cdot \Pr(A) + \Pr(H_1 \cap H_2|B) \cdot \Pr(B)}{\Pr(H_1)} = \frac{\Pr(H_1|A) \cdot \Pr(H_2|A) \cdot \Pr(A) + \Pr(H_1|B) \cdot \Pr(H_2|B) \cdot \Pr(B)}{\Pr(H_1)} = \frac{0.6 \cdot 0.6 \cdot 0.4 + 0.4 \cdot 0.4 \cdot 0.6}{0.48} = 0.5.$$

For now, let us disregard the cost x of testing the coin, and calculate the expected net payoff of the game without that additional fee. If Mr. Beige's coin test results in tails, his payoff for the game will be zero. Thus the expected net payoff for the case when he tests the coin (disregarding the fee of x dollars) is

$$-\$20 + \$20 \cdot \Pr(H_1^C) + \$250 \cdot \Pr(H_1) \cdot \Pr(H_2|H_1) - \$200 \cdot \Pr(H_1) \cdot \Pr(T_2|H_1) =$$

= -\\$20 + \\$20 \cdot 0.52 + \\$250 \cdot 0.48 \cdot 0.50 - \\$200 \cdot 0.48 \cdot 0.50 = \\$2.40.

This means that Mr. Beige's expected payoff changes from -\$4 to \$2.40 as a result of testing the coin. In order for him to be indifferent between the two choices, the additional

fee should be set at the difference of these two amounts, equal to his gain in the expected payoff of the game, i.e., \$6.40. Answer A.

Posted November 6, 2008 The first sentence of Problem 17 in Practice Examination 17 should be:

You are given a discrete random variable N such that its only possible values are 0, 1, 2,

3, 4, and 5. **instead of**

You are given a discrete random variable N such that its only possible values are 0, 1, 2, 3, 4, and 5, and $f_N(n+1) > f_N(n)$ for n = 0, 1, 2, 3, 4.

Posted November 4, 2008

 The answer choices in Problem 18, Practice Examination 16, should be:

 A. 0.3333
 B. 0.4875
 C. 0.6075
 D. 1.3333
 E. 2.1251

 They were mislabeled as A, B, B, C, E.

Posted October 5, 2008

Solution of Problem 29 in Practice Examination 18 should be:

Let W be the unconditional reimbursement amount, and let Y be the reimbursement, given that the reimbursement is positive. We have

$$W = \begin{cases} 0, & X \le 20, \\ X - 20 & 20 < X \le 120, \\ (120 - 20) + 0.5 \cdot (X - 120) = 40 + 0.5X & X > 120. \end{cases}$$

But Y = (W|X > 20), so that

$$Y = \begin{cases} X - 20 & X \le 120, \text{ given that } X > 20, \\ 40 + 0.5X & X > 120, \text{ given that } X > 20. \end{cases}$$

Note that when $X \le 120$, $X - 20 \le 100 < 115$. Also, X has exponential distribution, with memoryless property. We conclude that

$$G(115) = \Pr(Y \le 115) =$$

$$= \Pr(\{X \le 120 | X > 20\} \cup \{\{40 + 0.5X \le 115\} \cap \{X > 120\} | X > 20\}) =$$

$$= \frac{\Pr(20 < X \le 120)}{\Pr(X > 20)} + \Pr(\{120 < X \le 150 | X > 20\}) =$$

$$= \frac{\Pr(20 < X \le 120)}{\Pr(X > 20)} + \frac{\Pr(120 < X \le 150)}{\Pr(X > 20)} = \frac{\Pr(20 < X \le 150)}{\Pr(X > 20)}$$

$$= \Pr(X \le 130) = F_X(130) = 1 - e^{-\frac{130}{100}} \approx 0.727.$$

Answer B.

Posted September 1, 2008 In the solution of Problem 16 in Practice Examination 2, the left-hand side of the second formula should be $E(X^2)$, not E(X).

Posted August 5, 2008

In the solution of Problem 10 in Practice Examination 8, the phrase Then the total number of claims among 10 less than \$1,050 follows the binomial distribution with probability of success p = 0.8413should be Then the total number of claims among 10 less than \$1,050 follows the binomial distribution with probability of success p = 0.5793

Posted July 24, 2008

In the solution of Problem 6 of Practice Examination 7, the phrase

Therefore, recalling that for a discrete random variable, whose only possible values are possible integers,

should be

Therefore, recalling that for a discrete random variable, whose only possible values are positive integers,

Posted July 23, 2008 In the solution of Problem 5 in Practice Examination 3 the expression $Var(X_2) = 40,000$ should be $Var(X_2) = 250,000$.

Posted July 11, 2008 The formula at end of the third sentence of the solution of Problem 8 in Practice Examination 11 should be $\frac{0.26}{9} \approx 2.8889\%$ instead of $\frac{1.26}{9} \approx 2.89\%$.

Posted June 28, 2008 In the solution of Problem 18, Practice Examination 15, the sentence In the table used for Course P/1, $\Phi(1.39) = 0.9177$ and $\Phi(1.39) = 0.9192$. should be From the standard normal distribution table, $\Phi(1.39) = 0.9177$ and $\Phi(1.40) = 0.9192$.

Posted June 25, 2008

In Practice Examination 14, Problem 13, answer choice C should be 20762, and the final steps of its calculation should be:

In the table of the standard normal distribution we find $\Phi(0.67) = 0.7486$ and $\Phi(0.68) = 0.7517$. Using linear interpolation, we obtain the 75-th percentile of the standard normal distribution to be

$$z_{0.75} \approx 0.67 + \frac{0.75 - 0.7486}{0.7517 - 0.7486} \cdot (0.68 - 0.67) \approx 0.6745.$$

The maximum amount paid is the 75-th percentile of the distribution of (Y|X = 1200), and that is calculated as

 $20280 + z_{0.75} \cdot \sqrt{510000} \approx 20280 + 0.6745 \cdot \sqrt{510000} \approx 20761.6893.$ Answer C.

Posted June 15, 2008

Formula on page 49 should be

$$s_{T}(t) = \Pr(T > t) = \Pr(\min(T_{1}, T_{2}, \dots, T_{n}) > t) =$$

= $\Pr(\{T_{1} > t\} \cap \{T_{2} > t\} \cap \dots \cap \{T_{n} > t\}) = e^{-\lambda_{1}t} \cdot e^{-\lambda_{2}t} \cdot \dots e^{-\lambda_{n}t} = e^{-(\lambda_{1} + \dots + \lambda_{n})t}.$

instead of

$$s_{T}(t) = \Pr(T > t) = \Pr(\min(T_{1}, T_{2}, \dots, T_{n}) > t) =$$

= $\Pr(\{T_{1} > t\} \cap \{T_{2} > t\} \cap \dots \cap \{T_{n} > t\}) = e^{-\lambda_{1}t} \cdot e^{-\lambda_{1}t} \cdot \dots e^{-\lambda_{n}t} = e^{-(\lambda_{1} + \dots + \lambda_{n})t}$

Posted May 30, 2008

In the statement of Problem No. 1, Practice Examination 18, the expression

(iii) Exactly 40% of policyholders file fewer than claims during a given year. **should be:**

(iii) Exactly 40% of policyholders file fewer than two claims during a given year.

Posted May 30, 2008 In Problem No. 20, Practice Examination 18, the answer choice should be **C**.